

92.2.8. *Partitioned Regression with Rank-Deficient Regressions*, proposed by Peter C.B. Phillips. Let the m -vectors x_t ($t = 1, \dots, n$) be i.i.d. $N(0, \Sigma)$ and suppose that x_t is partitioned into subvectors of dimension m_0 , m_1 , and m_2 with $m = m_0 + m_1 + m_2$ as $x_t' = (x_{0t}', x_{1t}', x_{2t}')$. The covariance matrix Σ is partitioned into blocks as $\Sigma = (\Sigma_{ij})(i, j = 0, 1, 2)$, conformably with x_t . The matrices Σ, Σ_{ii} ($i = 0, 1, 2$) are all nonsingular.

(a) Show that the regression of x_{0t} on x_{1t} and x_{2t} can be written in the form

$$x_{0t} = \Sigma_{01} \Sigma_{11}^{-1} x_{1t} + \Sigma_{02 \cdot 1} \Sigma_{22 \cdot 1}^{-1} x_{2 \cdot 1t} + x_{0 \cdot 12t}, \quad (1)$$

where we use the notation,

$$\Sigma_{ij \cdot k} = \Sigma_{ij} - \Sigma_{ik} \Sigma_{kk}^{-1} \Sigma_{kj},$$

$$x_{i \cdot jt} = x_{it} - \Sigma_{ij} \Sigma_{jj}^{-1} x_{jt}.$$

The residual $x_{0 \cdot 12t}$ in (1) is independent of x_{1t} and x_{2t} and

$$x_{0 \cdot 12t} \equiv N(0, \Sigma_{00 \cdot 12}),$$

where

$$\Sigma_{00 \cdot 12} = \Sigma_{00} - \Sigma_{01} \Sigma_{11}^{-1} \Sigma_{10} - \Sigma_{02 \cdot 1} \Sigma_{22 \cdot 1}^{-1} \Sigma_{20 \cdot 1},$$

and “ \equiv ” signifies equivalence in distribution.

(b) Write the OLS regression equation of x_{0t} on x_{1t} and x_{2t} , as

$$x_{0t} = \hat{B}_1 x_{1t} + \hat{B}_2 x_{2t} + \hat{x}_{0 \cdot 12t}.$$

(i) Show that as $n \rightarrow \infty$

$$\hat{B}_1 \rightarrow_p B_1 = \Sigma_{01 \cdot 2} \Sigma_{11 \cdot 2}^{-1} = \Sigma_{01} \Sigma_{11}^{-1} - \Sigma_{02 \cdot 1} \Sigma_{22 \cdot 1}^{-1} \Sigma_{21} \Sigma_{11}^{-1},$$

$$\hat{B}_2 \rightarrow_p B_2 = \Sigma_{02 \cdot 1} \Sigma_{22 \cdot 1}^{-1} = \Sigma_{02} \Sigma_{22}^{-1} - \Sigma_{01 \cdot 2} \Sigma_{11 \cdot 2}^{-1} \Sigma_{12} \Sigma_{22}^{-1}.$$

(ii) Find the limit distribution as $n \rightarrow \infty$ of

$$\sqrt{n}[\hat{B}_1 - B_1, \hat{B}_2 - B_2]$$

and show that the limiting covariance between the matrices in the partition above is

$$-\Sigma_{00 \cdot 12} \otimes \Sigma_{11 \cdot 2}^{-1} \Sigma_{12} \Sigma_{22}^{-1} = -\Sigma_{00 \cdot 12} \otimes \Sigma_{11}^{-1} \Sigma_{12} \Sigma_{22}^{-1}.$$

Verify the equality above.

(c) Now suppose that the matrices Σ_{11} and Σ_{22} are singular.

(i) Show that the regression of x_{0t} on x_{1t} and x_{2t} can be written in the form

$$x_{0t} = B_1 x_{1t} + B_2 x_{2t} + x_{0 \cdot 12t}, \tag{1'}$$

where $B_1 = \Sigma_{01} \Sigma_{11}^- - \Sigma_{02 \cdot 1} \Sigma_{22 \cdot 1}^- \Sigma_{21} \Sigma_{11}^-$, $B_2 = \Sigma_{02 \cdot 1} \Sigma_{22 \cdot 1}^-$, $\Sigma_{22 \cdot 1}^- = \Sigma_{22} - \Sigma_{21} \Sigma_{11}^- \Sigma_{12}$, $\Sigma_{02 \cdot 1} = \Sigma_{02} - \Sigma_{01} \Sigma_{11}^- \Sigma_{12}$ and any choice of generalized inverses may be employed in these formulae. In (1') $x_{0 \cdot 12t} \equiv$ i.i.d. $N(0, \Sigma_{00 \cdot 12})$ and $\Sigma_{00 \cdot 12} = \Sigma_{00} - \Sigma_{01} \Sigma_{11}^- \Sigma_{10} - \Sigma_{02 \cdot 1} \Sigma_{22 \cdot 1}^- \Sigma_{20 \cdot 1}$, again for any choice of generalized inverse.

(ii) Are the coefficient matrices

$$B_1 = \Sigma_{01} \Sigma_{11}^- - \Sigma_{02 \cdot 1} \Sigma_{22 \cdot 1}^- \Sigma_{21} \Sigma_{11}^-,$$

and

$$B_2 = \Sigma_{02 \cdot 1} \Sigma_{22 \cdot 1}^-,$$

in (1') identified? If not, why not? Find matrices that are observationally equivalent to B_1 and B_2 . What asymptotically estimable functions of these coefficient matrices exist (i.e., what functions of the coefficient matrices B_1 and B_2 can be consistently estimated)?

(iii) Show that least-squares regression of x_{0t} on x_{1t} and x_{2t} leads to coefficient matrices \hat{B}_1 and \hat{B}_2 whose form is as follows:

$$\hat{B}_1 = X_0' X_1 (X_1' X_1)^- - X_0' Q_1 X_2 (X_2' Q_1 X_2)^- X_2' X_1 (X_1' X_1)^-,$$

and

$$\hat{B}_2 = X_0' Q_1 X_2 (X_2' Q_1 X_2)^{-},$$

where $Q_1 = I - X_1 (X_1' X_1)^{-} X_1'$ and any choice of generalized inverse is used. In these formulae the matrices $X_i' = [x_{i1}, \dots, x_{in}]$ are the data matrices of the variables x_{it} ($i = 0, 1, 2$).

- (iv) Find consistent estimators of the asymptotically estimable functions mentioned in part (ii) that are based on \hat{B}_1 and \hat{B}_2 as given in part (iii).