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I have for some years specialized in Open Book Problem Type examinations in Econometrics and enclose some recent material of mine which spans our current Graduate Econometrics Sequence here at Yale. It is organized as follows:

- (1) Graduate Introduction to Econometrics: Final Examination 1980
- (2) Graduate Econometrics: Final Examination 1980
- (3) Graduate Econometrics: Mid-term Examination 1981
- (4) Graduate Advanced Econometrics Problem Sequence 1980

All of this material involves new problems, ranging from introductory to advanced work. As such, it should reflect well the current style and content of our Econometrics Sequence at Yale.

One final point occurs to me. I may well wish to use-some of this material in later years myself in my Exercises in Econometrics series (published by Philip Allan and Ballinger).

Peter C. B. Phillips March 19, 1981

## YALE UNIVERSITY

## GRADUATE INTRODUCTION TO ECONOMETRICS

# FINAL EXAMINATION

Spring 1980

Author: P.C.B. Phillips

Time allowed:

THREE Hours.

Instructions:

Answer THREE questions.

This is an OPEN BOOK examination. ANY reference material allowed.

#### 1. In the model

$$y_t = \beta_1 x_{1t} + \beta_2 x_{2t} + \beta_3 x_{3t} + u_t$$

the  $x_{it}$  are non-random and  $u_t$  is a serially independent random disturbance with  $E(u_t)=0$  and  $E(u_t^2)=\sigma^2$  for all t. The coefficients in the model are known to be related to a more basic economic parameter  $\alpha$  according to the equations

$$\beta_1 + \beta_2 = \alpha$$

$$\beta_1 + \beta_3 = -\alpha$$
.

Given the following sample second moment matrix (obtained from a sample of 100 observations)

- (a) Find the best linear unbiased estimate of  $\alpha$  .
- (b) Without carrying out the calculations, show how to estimate the variance of  $\ \alpha$  .
- 2. In the model

$$y_1 = \beta_1 + u_1$$
 $y_2 = u_2$ 
 $y_3 = \beta_2 + u_3$ 
 $y_4 = u_4$ 

the  $y_t$  (t = 1, ...,4) are observable economic variables,  $\beta_1$  and  $\beta_2$  are unknown parameters and the  $u_t$  (t = 1, ...,4) are random disturbances with zero means and covariance matrix

$$\begin{bmatrix} 4 & 2 & 1 & \frac{1}{4} \\ 2 & 4 & 2 & 1 \\ 1 & 2 & 4 & 2 \\ \frac{1}{4} & 1 & 2 & 4 \end{bmatrix}$$

where  $\sigma^2 > 0$  is an unknown parameter.

- (a) Suggest a model for the random disturbances that is consistent with the stated properties of  $u_1,\ u_2,\ u_3,\ u_4$  .
- (b) Using only observations  $y_1$  and  $y_2$  find the best linear unbiased estimate of the parameter  $\beta_1$  .
- (c) If it is known that  $\beta_2 = 0$  , and all observations  $\gamma_1$ ,  $\gamma_2$ ,  $\gamma_3$ ,  $\gamma_4$  are available, find the best linear unbiased estimate of  $\beta_1$ .

- (d) Assuming that all observations  $y_1, y_2, y_3, y_4$  are again available but that  $\beta_2$  is unknown, find the best linear unbiased estimates of  $\beta_1$  and  $\beta_2$ .
- (e) Comment on your results in (b), (c) and (d).
- An investigator studying the demand for oranges specifies the relationship

$$Q_t = a_0 p_t^{a_1} q_t^{a_2} y_t^{a_3} e^{u_t}$$

where

 $Q_t$  = purchases of oranges in period t

 $p_t = price of oranges in period t$ 

 $q_{t}$  = price index of other fruit in period t

y, = nominal income in period t

u<sub>t</sub> = random disturbance.

Using T=50 time series observations, the following regression is estimated by ordinary least squares

(Standard errors are given in parentheses.)

Noting that the own price elasticity has the wrong sign in this regression and that neither price coefficient is significantly different

from zero, the investigator attempts to improve the results by a constrained regression which yields

$$lnQ_t = const - 0.24ln \quad \left(\frac{P_t}{q_t}\right) + 0.52ln \quad \left(\frac{Y_t}{q_t}\right)$$

$$R^2 = 0.73$$
 Dw = 0.55

The investigator is happy with the outcome of his regression because it embodies a constraint that has a meaningful economic interpretation and the estimates obtained are of the right sign and also seem to be statistically significant. However, a colleague points out to him that the Durbin Watson statistic is rather low and suggests that he run a new regression including  $\ln Q_{t-1}$  as a regressor. The results are

$$\ln Q_{t} = \text{const} - \underset{(0.05)}{\text{0.17\&n}} \quad \left[ \frac{P_{t}}{q_{t}} \right] \quad + \underset{(0.28)}{\text{0.34\&n}} \quad \left[ \frac{y_{t}}{q_{t}} \right]$$

$$R^2 = 0.89$$

Both investigators are satisfied with the outcome of this final regression.

DW = 1.8

Critically appraise the approach to model specification, estimation and diagnostic checking taken in this empirical investigation, indicating areas of strength and weakness in the exercise. Try to justify your criticisms and your support in the light of econometric theory and the empirical evidence reported.

## 4. In the model

$$y_{1t} = \alpha x_{1t} + u_{1t}$$

$$y_{2t} = \beta x_{2t} + u_{2t}$$

The  $\mathbf{x}_{it}$  are non-random exogenous variables and the  $\mathbf{u}_{it}$  are serially independent random disturbances that are normally distributed with zero means and second moments

$$E(u_{1t}^2) = 1$$
 ,  $E(u_{2t}^2) = 2$  ,  $E(u_{1t}^2) = 1$ 

for all values of t . The sample second moment matrix below was calculated from 20 sample observations:

- (a) Find the best linear unbiased estimates of the parameters  $\,\alpha\,$  and  $\,\beta\,$  in the above model.
- (b) Test the null hypothesis

against the alternative

5. An investigator has specified the following two models and proposes to use them in some empirical work with macroeconomic time series data:

$$c_t = \alpha_1 y_t + \alpha_2 m_{t-1} + u_{1t}$$

$$i_t = \beta_1 y_t + \beta_2 r_t + u_{2t}$$

$$y_t = c_t + i_t$$
(Model 1)

jointly dependent variables:  $c_t$ ,  $i_t$ ,  $y_t$  predetermined variables:  $r_t$ ,  $m_{t-1}$ 

$$m_{t} = \gamma_{1} r_{t} + \gamma_{2} m_{t-1} + v_{1t}$$

$$r_{t} = \delta_{1} m_{t} + \delta_{2} m_{t-1} + \delta_{3} y_{t} + v_{2t}$$
(Model 2)

jointly dependent variables:  $\mathbf{m_t}$ ,  $\mathbf{r_t}$  predetermined variables:  $\mathbf{m_{t-1}}$ ,  $\mathbf{y_t}$  .

- (a) Assess the identifiability of the parameters that appear as coefficients in the above two models (treating the two models separately).
- (b) Obtain the reduced form equation for  $y_t$  in model (1) and the reduced form equation for  $r_t$  in model (2).
- (c) Assess the identifiability of the two-equation model comprising the reduced form equation for  $y_t$  in model (1) (an IS curve) and the reduced form equation for  $r_t$  in model (2) (an IM curve).
- 6. The two equations of a competitive model of the meat market are:

Demand: 
$$y_{1t} = a_1 y_{2t} + a_2 x_{1t} + u_{1t}$$

Supply: 
$$y_{1t} = b_1 y_{2t} + b_2 x_{2t} + b_3 x_{3t} + u_{2t}$$

where

 $\gamma_1$  = meat consumption per capita (dollars p.a.) and

 $y_2$  = meat price (1935 = 100) are jointly dependent variables;

 $x_1 = disposable income per capita (dollars p.a.),$ 

 $\rm m_2$  = unit cost of meat processing (1935 = 100) and

 $x_3$  = cost of agricultural production (1935 = 100) are exogenous variables;

 $^{\mathrm{u}}$ lt'  $^{\mathrm{u}}$ 2t are stochastic disturbances.

(a) Estimate the parameters  $\ {\bf a}_1$  and  $\ {\bf a}_2$  by two stage least squares, and test the hypothesis

$$H_0: a_2 = 0$$
 against  $H_1: a_2 \neq 0$ 

given the following sample second moment matrix (based on 36 observations):

	Y <sub>1</sub>	У <sub>2</sub>	×1	* <sub>2</sub>	<b>x</b> 3
У1	10	0	1	0	-1
У2	0	70	-1	~1	0
×1	1	-1	1	0	0
× <sub>2</sub>	0	-1	0	1	0
×,	-1	0	0	0	1

## YALE UNIVERSITY

#### GRADUATE ECONOMETRICS

## FINAL EXAMINATION

Spring 1980

Author: P.C.B. Phillips

Time allowed:

THREE Hours.

Instructions:

Answer THREE questions.

This is an OPEN BOOK examination. ANY reference material allowed.

1. Two economic variables  $\,\eta_{\,{\mbox{\scriptsize t}}}\,\,$  and  $\,\,\xi_{\,{\mbox{\scriptsize t}}}\,\,$  are assumed to be related by the equation

$$n_t = \alpha \xi_t$$
  $(t = 1, ..., T)$ 

where  $\alpha$  is an unknown parameter. Observable variables  $\gamma_t$  and  $x_t$  are known to be related to  $\eta_t$  and  $\xi_t$  according to the equations

$$y_t = \eta_t + u_t$$

$$x_t = \xi_t + v_t$$

where  $u_t$  and  $v_t$  are serially independent random disturbances which are distributed independently of t and of  $n_g$  and  $\xi_g$  (s = 1, ..., T) with zero means and higher moments given by

$$\begin{split} \mathbb{E}(\mathbf{u}_t^2) &= \sigma_\mathbf{u}^2 \ , \ \mathbb{E}(\mathbf{u}_t^2) = \sigma_\mathbf{u}^2 \ , \ \mathbb{E}(\mathbf{u}_t^3) = 0 \ , \ \mathbb{E}(\mathbf{u}_t^4) = 3\sigma_\mathbf{u}^4 \ , \\ \mathbb{E}(\mathbf{u}_t v_t) &= 0 \end{split}$$

for all t .

Given that  $\xi_{t} = m(1+r^{t})$  , where  $m \geq 0$  and  $0 \leq r \leq 1$  are constants:

- (a) Find the limit in probability as  $T \gg \infty$  of  $\alpha^{\bigstar}$  , the estimator of  $\alpha$  obtained from an ordinary least squares regression of  $y_{t}$  on  $x_{t}$  ;
- (b) for the special case in which  $\sigma_u^2 = 0$  so that  $\gamma_t = \eta_t$  with probability one, prove that  $\sqrt{T}(\alpha^* \text{plim } \alpha^*)$  has a limiting normal distribution and find the variance of this distribution.
- 2. In the model

$$y_1 = 2\beta_1 + \beta_2 + \beta_3 + u_1$$
  
 $y_2 = \beta_1 + \beta_2 + u_2$   
 $y_3 = \beta_1 + \beta_3 + u_3$ 

The  $y_t(t=1, 2, 3)$  are observable economic variables,  $\beta_1$ ,  $\beta_2$ ,  $\beta_3$  are unknown parameters and the  $u_t(t=1, 2, 3)$  are random disturbances distributed with zero means, constant variance  $E(u_t^2) = \sigma^2$  for all t and  $E(u_t, u_s) = 0$  for t = 1.

- (a) Show that  $\beta_1$   $\beta_2$  +  $2\beta_3$  is an estimable function.
- (b) Find the best linear unbiased estimate of  $~\beta_1$   $\beta_2$  +  $2\beta_3~$  .
- (c) Find an unbiased estimate of the variance  $\ensuremath{\sigma^2}$  .

[Hint: If M is an n x n positive semi definite matrix of rank n-1 and d is a non-zero vector for which Md = 0 then a generalized inverse of M is given by the matrix  $[M + dd^{\dagger}]^{-1}$ .

3. An investigator studying the demand for money specifies the relationship

$$m_{\rm t} = \alpha_0 + \alpha_1 y_{\rm t} + \alpha_2 x_{\rm t} + u_{\rm t}$$

where

 $m_t$  = volume of money in period t  $y_t$  = nominal income in period t  $r_t$  < rate of interest in period t  $u_t$  = random disturbance.

The relationship is estimated by ordinary least squares (OLS) from a time series of T = 50 observations, for which the correlations between  $\mathbf{m_t}$  and  $\mathbf{t}$  and  $\mathbf{y_t}$  and  $\mathbf{t}$  are 0.995 and 0.980 respectively. The result is

$$m_t = const + 2.11 y_t - 4.01 r_t$$

$$(0.05) t (30.11)$$

$$R^2 = 0.99 DW = 0.51$$

where standard errors are in parentheses. Since the coefficient of  $\mathbf{r}_{t}$  is insignificant he drops this variable but because of the low DW (Durbin-Watson) statistic he decides to include  $\mathbf{m}_{t-1}$  as a regressor. Reestimation of the equation by OLS yields

$$m_t = const + 0.9 y_t + 0.7 m_{t-1}$$
  
 $(0.04) (0.14)$   
 $R^2 = 0.996$   $DW = 2.3$ 

which the investigator accepts as a satisfactory relationship.

Comment on the approach to model specification, estimation and diagnostic checking taken in this empirical investigation. Support your comments by evaluating the empirical results reported and in the light of your knowledge of econometric theory.

4. In the two equation regression model

$$\begin{bmatrix} y_1 \\ y_2 \end{bmatrix} = \begin{bmatrix} x_1 & 0 \\ 0 & x_2 \end{bmatrix} \begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} + \begin{bmatrix} u_1 \\ u_2 \end{bmatrix}$$

 $y_{\underline{i}}$  (i = 1, 2) is a vector of T observations on the i'th dependent variable,  $x_{\underline{i}}$  is a matrix of observations on  $m_{\underline{i}}$  non-random independent variables, the  $\beta_{\underline{i}}$  are vectors of unknown coefficients and the  $u_{\underline{i}}$  are vectors of disturbances for which

$$E(u_1) = E(u_2) = 0$$
  
 $E(u_1u_1) = \sigma_{11}I_T$ ,  $E(u_2u_2) = \sigma_{22}I_T$ ,  $E(u_1u_2) = \sigma_{12}I_T$ .

It is assumed that the matrix

$$\sum = \begin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{12} & \sigma_{22} \end{bmatrix}$$

is positive definite and known.

(a) If  $\hat{\beta}_1$  and  $\hat{\beta}_2$  denote the Zellner seemingly unrelated regression estimator of  $\beta_1$  and  $\beta_2$  respectively, show that the covariance matrices of these estimators are given by

$$\sigma_{11}^{(1-\rho^2)} \left[ x_1^{\prime} x_1^{\prime} - \rho^2 x_1^{\prime} x_2^{\prime} (x_2^{\prime} x_2^{\prime})^{-1} x_2^{\prime} x_1^{\prime} \right]^{-1}$$

and

$$\sigma_{22}^{-}(1-\rho^2) \ \left[ x_2^{'} x_2^{} - \rho^2 \ x_2^{'} \ x_1^{-} (x_1^{'} x_1)^{-1} x_1^{'} x_2^{} \right]^{-1} \ ,$$

respectively, where  $\rho = \sigma_{12}/(\sigma_{11}\sigma_{22})^{\frac{1}{2}}$  .

- (b) Hence, prove the result that the single equation least squares estimators of  $\beta_1$  and  $\beta_2$  are as efficient as  $\hat{\beta}_1$  and  $\hat{\beta}_2$   $\underline{\text{if and only if either } \rho = 0 \text{ or the range spaces of } x_1 \text{ and } x_2$  are equivalent.
- 5. In the simultaneous equations model

$$By_t + Cx_t = u_t$$

 $\mathbf{y_t}$  is a vector of observations in period t on n endogenous variables,  $\mathbf{x_t}$  is a vector of observations in period t on m exogenous variables,  $\mathbf{u_t}$  is a vector of serially independent random disturbances distributed with zero means and non-singular covariance matrix and  $\mathbf{B}(\mathbf{n}\mathbf{x}\mathbf{n})$  and  $\mathbf{C}(\mathbf{n}\mathbf{x}\mathbf{m})$  are matrices of coefficients.  $\mathbf{B}$  is non-singular.

Let b' and c' denote the first rows of the matrices B and C respectively. It is known that the elements of b and c satisfy the following restrictions (which include the normalization rule)

where  $\Phi = \begin{bmatrix} \phi_1 : \phi_2 \end{bmatrix}$  is a known rx(n+m) matrix and  $\phi$  is a known vector.

- (a) Show that if the first equation is identified then
  - (i) There exists a vector  $x \neq 0$  for which

$$\left[\phi_1 - \phi_2 A^{\dagger}\right] \times = \phi$$

(ii) rank 
$$\left[ \phi_1 \sim \phi_2 A' \right] = n$$

where  $A = B^{-1} C$  is the reduced form coefficient matrix.

(b) For the special case where n=2 and the first equation is

$$b_{11}y_{1t} + b_{12}y_{2t} + c_{11}x_{1t} + c_{12}x_{2t} + c_{13}x_{3t}$$
  
+  $c_{14}x_{4t} = u_{1t}$ 

find the restrictions on the reduced form matrix A implied by the restrictions

$$b_{12} = c_{11}$$
  $c_{12} = c_{13}$   $c_{11} + c_{13} = 1$ 

with the normalization  $b_{11} = 1$  .

6. In the model

$$y_{it} = \alpha_1 y_{2t} + \alpha_2 x_{1t} + u_{1t}$$
  
 $y_{2t} = \beta_1 y_{1t} + \beta_2 x_{2t} + u_{2t}$ 

the  $y_{it}$  are endogenous variables and the  $x_{it}$  are exogenous variables which we assume to be non-random and whose moment matrix we assume to have a finite positive definite limit as the sample size tends to infinity. The random disturbances  $u_{it}$  are scrially independent and are distributed with zero means and second moments

$$E(u_{1t}^2) = 1$$
 ,  $E(u_{1t}^2u_{2t}) = 1$  ,  $E(u_{2t}^2) = 2$ 

for all t .

Find consistent estimates of the parameters  $\alpha_1$ ,  $\alpha_2$ ,  $\beta_1$  and  $\beta_2$  given the following sample second moment matrix (based on T = 100 observations) and test the null hypothesis

$$((a_{ij}+a_{j})+\phi_{j})\mapsto (a_{ij}+a_{j})$$

operated the absolution

$$\mathcal{B}_{1}^{-1}: \mathcal{A}_{1}^{-1} \times \mathcal{B}_{2}^{-1} \times \mathcal{B}_{3}^{-1} \times \mathcal{B}_{3}^{-1} \times \mathcal{B}_{3}^{-1} = \emptyset$$

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## YALE UNIVERSITY

#### GRADUATE ECONOMETRICS

## MID TERM EXAMINATION

Winter 1981:

Author: P.C.B. Phillips

Time allowed: ONE Hour and a HALF.

Instructions: Answer ONE question.

This is an OPEN BOOK examination. ANY reference material allowed.

 An investigator assumes that two observable economic variables are related by the equation

$$y_t = \alpha x_t + u_t$$
 (t = 1, ...,T) (1)

where  $\alpha$  is an unknown parameter and the  $~u_{t}^{}$  (t = 1, ..., T) are assumed to be independent and identically distributed (i.i.d.)  $N~(o,\sigma_{u}^{2})~.~In~fact,~the~investigator's~equation~(1)~is~misspecified~and~the~true~relationship~is~given~by$ 

$$y_t = \alpha x_t + \beta z_t + \varepsilon_t$$
  $(t = 1, \dots, T)$  (2)

which involves a third observable economic variable  $z_{t}$ , an additional unknown coefficient  $\beta$  and the  $\varepsilon_{t}$  are i.i.d. N  $(o,\sigma_{\varepsilon}^{2})$ . In the true relationship (2) the  $x_{t}$  are i.i.d. N  $(m_{x},\sigma_{x}^{2})$ , the variable  $z_{t}$  is generated by the equation

$$z_{t} = \gamma w_{t} + \eta_{t} \qquad (t = 1, \dots, T)$$
 (3)

in which  $\gamma$  is an unknown parameter, the  $w_{t}$  are i.i.d. N  $(n_{w},\sigma_{w}^{2})$  and the  $n_{t}$  are i.i.d. N  $(o,\sigma_{\eta}^{2})$  . The random variables  $\epsilon_{t},n_{t},x_{t}$  and

- $\mathbf{w}_{\mathbf{t}}$  in equations (2) and (3) are statistically independent.
- (a) Find the limit in probability as  $T \to \infty$  of  $\alpha^*$  , the least squares estimator of  $\alpha$  in the investigator's equation (1).
- (b) Find the limiting distribution as  $T \to \infty$  of  $\sqrt{T}$  ( $\alpha*$  plim  $\alpha*$ ) and derive explicitly a formula for the variance of this distribution.
- 2. The observable nxl random vectors  $\mathbf{y_1}, \dots, \mathbf{y_T}$  and the non-random mxl vectors  $\mathbf{x_1}, \dots, \mathbf{x_T}$  satisfy the system

$$y_t = Ax_t + u_t$$
 (t = 1, ..., T)

where A is an nxm matrix of unknown coefficients and the  $u_t$   $(t=1,\ldots,T)$  are serially independent random vectors which are identically distributed as  $N(o,\sigma^2\Sigma)$  where  $\Sigma$  is a known positive definite nxn matrix and  $\sigma^2$  is an unknown scalar. It is assumed that T>m and that the data matrix  $X'=\left[x_1,\ldots,x_T\right]$  has full rank.

- (a) Show that the best linear unbiased estimator of the coefficient matrix A is given by the ordinary least squares estimator  $A^* = Y'X(X'X)^{-1} \text{ where } Y' = \begin{bmatrix} y_1, \dots, y_T \end{bmatrix}.$
- (b) Construct a statistical test of the hypothesis that the rows of the matrix A are identical (i.e. that the coefficients of the explanatory variables are the same for all equations in the model). Find the exact sampling distribution of your test statistic and hence justify your testing procedure.

#### YALE UNIVERSITY

#### ADVANCED ECONOMETRICS

## PROBLEM SEQUENCE IN LINEAR AND NON-LINEAR ESTIMATION

Fall 1980

Author: P.C.B. Phillips

# 1. In the model

$$\begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} y_1 \\ -y_1 \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \end{bmatrix} \quad \text{or} \quad x = y + \varepsilon$$

the scalar  $~\gamma_1^{}~$  is unknown and the error vector  $~\epsilon~$  is known to have zero mean and covariance matrix

$$Q = \begin{bmatrix} \sigma^2 & \sigma^2 \\ & & \\ \sigma^2 & 2\sigma^2 \end{bmatrix} .$$

## (a) Draw a diagram to illustrate

- (1) the concentration ellipsoid of  $\epsilon$  ;
- (2) the linear subspace L in which the vector y is known to lie;
- (3) the linear subspace K which is the principal conjugate of L with respect to the concentration ellipsoid of  $\epsilon$  ;
- (4) the projection estimator of y in L based on the observation x, whose concentration ellipsoid is contained in that of every other unbiased linear estimator of y.
- (b) If  $x_1 = 10$  and  $x_2 = 1$  , calculate the best linear unbiased estimator of y .

2. In the linear model

$$\begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} = \begin{bmatrix} y_1 \\ y_1 \\ y_1 \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \varepsilon_3 \end{bmatrix} \quad \text{or} \quad x = y + \varepsilon$$

the mean of  $\epsilon$  is zero and its covariance matrix is given by

$$Q = \begin{bmatrix} \sigma^2 & 0 & 0 \\ 0 & 4\sigma^2 & 4\sigma^2 \\ 0 & 4\sigma^2 & 4\sigma^2 \end{bmatrix}$$

- (a) Find the linear transformation of  $\mathbb{R}^3$  into itself which takes the support of  $\varepsilon$  into the leading 2-dimensional subspace of  $\mathbb{R}^3$ , the concentration ellipsoid of  $\varepsilon$  into the unit circle in this subspace and the linear subspace L which contains the vector y into the first coordinate axis.
- (b) If H is the matrix of the transformation in (a), find the best linear unbiased estimator of Hy . Use this estimator to deduce the corresponding estimator of  $\gamma_1$  .
- (c) Show that the estimator of  $y_1$  obtained in (b) is identical to the estimator obtained from the projection on L along the principal conjugate subspace of L with respect to the concentration ellipsoid of  $\varepsilon$  in the original (untransformed) space.

## 3. In the linear model

$$\begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} y_1 \\ -y_1 \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \end{bmatrix} \quad \text{or } x + y + \varepsilon$$

the mean of  $\epsilon$  is zero and its covariance matrix is given by

$$Q = \begin{bmatrix} \sigma^2 & 2\sigma^2 \\ \\ 2\sigma^2 & 4\sigma^2 \end{bmatrix} .$$

- (a) Find the concentration ellipsoid of  $\,\epsilon\,$  .
- (b) Find the best linear unbiased estimator of  $\ y_1$  . Does your estimator have any other interesting properties?
- (c) Compare the variance of your estimator in (b) with that of the least squares estimator of  $\ y_1$  .
- 4. In the model

$$\begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} = \begin{bmatrix} y_1 \\ y_1 + y_2 \\ y_2 \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \varepsilon_3 \end{bmatrix} \quad \text{or } x = y + \varepsilon$$

the scalars  $~y_1^{}$  and  $~y_2^{}$  are unknown and the error vector  $~\epsilon~$  is known to have zero mean and covariance matrix

$$Q = \begin{bmatrix} 1 & 1 & 0 \\ 1 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}$$

- (a) Draw a figure to illustrate:
  - (1) the concentration ellipsoid of  $\epsilon$  ;
  - (2) the linear subspace  $\ L$  in which the mean vector  $\ y$  is known to lie;

- (3) the projection estimator of y in L based on the observation x , whose concentration ellipsoid is contained in that of every other unbiased linear estimator of y .
- (b) If the observation vector  $\ensuremath{\mathbf{x}}$  takes the value

$$x = \begin{bmatrix} 1 \\ 5 \\ 10 \end{bmatrix}$$

calculate the best linear unbiased estimator of y .

5. In the model

$$x_{1t} = a_{11}z_{1t} + a_{12}z_{2t} + u_{1t}$$
  
 $x_{2t} = a_{21}z_{1t} + a_{22}z_{2t} + u_{2t}$ 

the  $z_{it}$  are non-random and the  $u_{it}$  are serially independent random disturbances with zero means and covariance matrix

$$\Sigma = \begin{bmatrix} \sigma^2 & -\sigma^2 \\ & & \\ -\sigma^2 & \sigma^2 \end{bmatrix}$$

for all values of t .

(a) Find the best linear unbiased estimates of  $a_{11}$ ,  $a_{12}$ ,  $a_{21}$  and  $a_{22}$  in the above model, given the following matrix of sample second moments of the data.

- (b) If it is known that  $a_{12}=0$  , find the best linear unbiased estimates of  $a_{11}$  ,  $a_{21}$  and  $a_{22}$  .
- (c) Discuss the role of the singularity of the error covariance matrix  $\Sigma \quad \text{in your answers to (a) and (b).} \quad \text{Do we need a generalized inverse}$  of  $\Sigma$  to find the best linear unbiased estimates in part (b)?

#### 6. In the model

$$x_{1t} = \alpha z_{1t} + \beta z_{2t} + u_{1t}$$
 $x_{2t} = \alpha z_{2t} + u_{2t}$ 
 $x_{3t} = \beta z_{1t} + u_{3t}$ 

the  $\mathbf{z}_{\text{it}}$  are non-random and the  $\mathbf{u}_{\text{it}}$  are random disturbances with zero means and covariance matrix

$$\begin{bmatrix} 2\sigma^2 & \sigma^2 & \sigma^2 \\ \sigma^2 & \sigma^2 & 0 \\ \sigma^2 & 0 & \sigma^2 \end{bmatrix}$$

for each value of t . The  $\,u_{it}^{}$  are serially independent. Find the best linear unbiased estimates of  $\,\alpha\,$  and  $\,\beta\,$  in the above model given the following matrix of second moments based on a sample of size  $\,T\,$ :

- 7. The random n-vector x is assumed to have an unknown mean vector y, which is known to lie in a certain subspace L of dimension p in n dimensional Euclidean space, and a known covariance matrix Q of rank r < n. L is assumed to lie in the range space (or support) of Q.
  - (a) If  $y^*$  is the Gauss Markov (best linear unbiased) estimator of y, prove that  $y^*$  is the vector of L which minimizes the quantity  $(x-y)^*Q^-(x-y)$  for all y in L and any choice of generalized inverse  $Q^-$  of Q.
  - (b) Prove that  $y^*$  is identical to the ordinary least squares estimator of y if and only if K, the principal conjugate subspace of L with respect to the concentration ellipsoid of x is orthogonal to the subspace L.
  - (c) Hence, prove that  $y^*$  is identical to the ordinary least squares estimator of y if and only if the image of L under the transformation Q is identical to L.
- 8. In the model

$$\begin{bmatrix} x_{1t} \\ x_{2t} \end{bmatrix} = \begin{bmatrix} y_1 \\ y_2 \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \end{bmatrix} \quad \text{or} \quad x_t = y + \varepsilon_t \quad (t = 1, \dots, T)$$

it is known that  $y_2 = y_1^2$ , and the  $\varepsilon_{\mathsf{t}}$  are serially independent and normally distributed with zero mean vector and covariance matrix

$$\Omega = \begin{bmatrix} 2\sigma^2 & \sigma^2 \\ & & \\ \sigma^2 & \sigma^2 \end{bmatrix}$$

for each value of t .

(a) Given the following sample moments obtained from  $\, T = 100 \,$  observations

$$x_1$$
  $x_2$   
 $x_1$  9 -1  $\overline{x}_1 = 1.75$   $\overline{x}_2 = 1.75$   
 $x_2$  -1 9

find asymptotically efficient estimates of  $\mathbf{y}_1$  and  $\mathbf{y}_2$  .

- (b) Illustrate your estimates on a diagram.
- (c) How would you estimate  $\sigma^2$  ?
- 9. In the model

$$x_{t} = a_{1}z_{1t} + a_{2}z_{2t} + \varepsilon_{t}$$
  $t = 1, ..., T$ 

the  $z_{it}$  are non-random and the  $\varepsilon_t$  are normally distributed with  $E(\varepsilon_t)=0$ ,  $E(\varepsilon_t^2)=\sigma^2$  and  $E(\varepsilon_t\varepsilon_s)=0$ ,  $s\neq t$ .  $a_1$  and  $a_2$  satisfy the restriction  $a_1^2a_2=1$ .

(a) Explain how to find asymptotically efficient estimators  $\hat{a}_1$  and  $\hat{a}_2$  of  $a_1$  and  $a_2$  and compute the asymptotic variance of  $\sqrt{T}(a_1-a_2)$  when the true values of the parameters are  $c^2=a_1$ ,  $a_1=1$ ,  $a_2=1$  and

$$\lim_{T\to\infty}\frac{1}{T}\quad \sum_{t}\begin{bmatrix}z_{1t}&&z_{1t}z_{2t}\\\\z_{2t}z_{1t}&&z_{2t}\end{bmatrix}\quad =\quad \begin{bmatrix}1&1\\\\1&2\end{bmatrix}$$

(b) Compare your result with the asymptotic variance of  $\sqrt{T}(a_1^{\star}-a_2^{\star})$  where  $a_1^{\star}$  and  $a_2^{\star}$  denote the ordinary least squares estimators

of  $a_1$  and  $a_2$  .

10. In the model

$$x_t = ax_t z_t + bz_t + u_t$$

 $z_{t}^{}$  is non-random the  $~u_{t}^{}$  (t = 1, ... ,T) are independent, identically distributed random disturbances with zero mean and variance  $~\sigma^2~$  .

- (a) Explain how you would obtain consistent estimators of the parameters a and b in the above model by the method of instrumental variables (with instruments  $\mathbf{z}_{\mathsf{t}}$  and a constant). Indicate assumptions over and above those already given which will ensure that your estimators are consistent.
- (b) Find the covariance matrix of the limiting distribution of  $\sqrt{T}\,(\hat{a}-a\ ,\,\hat{b}-b) \quad \text{where} \quad \hat{a} \quad \text{and} \quad \hat{b} \quad \text{denote the instrumental}$  variable estimates of part (a) and

$$a = 1$$
,  $b = 2$ ,  $\sigma^2 = 1$ 

and

$$\lim_{T \to \infty} \frac{1}{T} \int_{t=1}^{T} z_{t} = 1 , \lim_{T \to \infty} \frac{1}{T} \int_{t=1}^{T} z_{t}^{2} = 2 , \lim_{T \to \infty} \frac{1}{T} \int_{t=1}^{T} \frac{z_{t}^{i}}{1 - z_{t}} = \frac{2^{i+1}}{2^{i} - 1}$$